

REMARKS ON OSTROWSKI'S AND HADAMARD'S INEQUALITY

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Starting from Taylor's formula with an integral form of remainder, we derive an inequality which, as special cases, yields Ostrowski's and Hadamard's inequality. We further present a two sided version of Ostrowski's inequality for monotone functions and apply the result in derivation of an inequality for tails of a probability distribution concentrated on an interval.

1. AN INEQUALITY FOR DIFFERENTIABLE FUNCTIONS

Let $C_b^{(n+1)}[a, b]$ be the set of real valued functions g defined on $[a, b]$, n times differentiable at every point and with $g^{(n+1)}(x)$ that exists everywhere on $[a, b]$ except possibly finitely many points, such that $g^{(n+1)}$ is RIEMANN integrable over $[a, b]$, hence bounded. For a $g \in C_b^{(n+1)}[a, b]$, apply TAYLOR's expansion of order n around an arbitrary $x \in (a, b)$, at points a and b respectively, to get

$$g(a) = g(x) + g'(x)(a-x) + \sum_{k=2}^n \frac{g^{(k)}(x)}{k!} (a-x)^k + \frac{1}{n!} \int_x^a (a-\tau)^n g^{(n+1)}(\tau) d\tau$$

$$g(b) = g(x) + g'(x)(b-x) + \sum_{k=2}^n \frac{g^{(k)}(x)}{k!} (b-x)^k + \frac{1}{n!} \int_x^b (b-\tau)^n g^{(n+1)}(\tau) d\tau,$$

where for $n = 1$, the sums are empty. By subtracting and dividing by $b - a$, we get

$$(1) \quad \frac{g(b) - g(a)}{b - a} - g'(x) = \sum_{k=2}^n \frac{g^{(k)}(x)}{k!} \cdot \frac{(b-x)^k - (a-x)^k}{b-a} + R_n,$$

where, after some work one can show that

$$(2) \quad R_n = \frac{(b-a)^n}{n!} \int_0^1 K(t) g^{(n+1)}(a + t(b-a)) dt$$

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and

$$(3) \quad K(t) = (-1)^n t^n I_{[0, (x-a)/(b-a)]}(t) + (1-t)^n I_{((x-a)/(b-a), 1]}(t).$$

Let $B_{n+1} = \sup_{x \in [a, b]} |g^{n+1}(x)|$. Then for an even n we have that $K(t) \geq 0$ and hence

$$(4) \quad \begin{aligned} |R_n| &\leq B_{n+1} \frac{(b-a)^n}{n!} \int_0^1 K(t) dt \\ &= \frac{(x-a)^{n+1} + (b-x)^{n+1}}{(b-a)(n+1)!} B_{n+1}. \end{aligned}$$

Note that if $g^{(n+1)}$ is of a constant sign on $[a, b]$ and n is even, then also R_n is of that sign.

For an odd n , the estimate (4) also holds; moreover, if $g^{(n+1)}(x) > 0$ for all $x \in [a, b]$ then from (2) and (3) we get

$$(5) \quad -\frac{(x-a)^{n+1}}{(b-a)(n+1)!} B_{n+1} \leq R_n \leq \frac{(b-x)^{n+1}}{(b-a)(n+1)!} B_{n+1},$$

and if $g^{(n+1)} < 0$ then

$$(6) \quad -\frac{(b-x)^{n+1}}{(b-a)(n+1)!} B_{n+1} \leq R_n \leq \frac{(x-a)^{n+1}}{(b-a)(n+1)!} B_{n+1},$$

Therefore, we have the following theorem.

Theorem 1. *Let $g \in C_b^{(n+1)}[a, b]$, $n \geq 1$. Further, let $B_{n+1} = \sup_{x \in [a, b]} |g^{(n+1)}(x)|$. For an arbitrary $x \in [a, b]$, define*

$$(7) \quad R_n(a, b, x; g) = \frac{g(b) - g(a)}{b-a} - g'(x) - \sum_{k=2}^n \frac{g^{(k)}(x)}{k!} \cdot \frac{(b-x)^k - (a-x)^k}{b-a},$$

where the sum is void for $n = 1$. Then the integral representation (2) holds and we have the following inequality

$$(8) \quad |R_n(a, b, x; g)| \leq \frac{(x-a)^{n+1} + (b-x)^{n+1}}{(b-a)(n+1)!} B_{n+1}.$$

If $g^{(2m+1)}(x) \geq 0$ for all $x \in [a, b]$, then $R_{2m}(a, b, x; g) \geq 0$. If $n = 2m - 1$, $m = 1, 2, \dots$ and $g^{2m}(x) \geq 0$ for all $x \in [a, b]$, then

$$(9) \quad -\frac{(x-a)^{2m}}{(b-a)(2m)!} B_{2m} \leq R_{2m-1}(a, b, x; g) \leq \frac{(b-x)^{2m}}{(b-a)(2m)!} B_{2m}.$$

The technique which was used here was also applied in [3] to obtain expressions for error terms in JENSEN's and some related inequalities.

In the next section we observe some special cases of Theorem 1.

2. SPECIAL CASES: OSTROWSKI'S AND HADAMARD'S INEQUALITIES

Let $f \in C_b^{(n)}[a, b]$ and let

$$(10) \quad g(x) = \int_a^x f(t) dt$$

Then the expression (7) reads

$$(11) \quad R_n(a, b, x; g) = \frac{1}{b-a} \int_a^b f(t) dt - f(x) - \sum_{k=2}^n \frac{f^{(k-1)}(x)}{k!} \cdot \frac{(b-x)^k - (a-x)^k}{b-a}$$

Let $M_n = B_{n+1} = \sup_{x \in [a, b]} |f^{(n)}(x)|$. Theorem 1 for $n = 1$ gives OSTROWSKI's inequality:

$$(12) \quad \left| \frac{1}{b-a} \int_a^b f(t) dt - f(x) \right| \leq \frac{(x-a)^2 + (b-x)^2}{2(b-a)} M_1.$$

The bound in (12) is usually written in the form [2]:

$$(b-a) \left(\frac{1}{4} + \frac{(x - (a+b)/2)^2}{(b-a)^2} \right) M_1.$$

If $f' \geq 0$ on $[a, b]$, then by Theorem 1 we have

$$(13) \quad -\frac{(x-a)^2}{2(b-a)} M_1 \leq \frac{1}{b-a} \int_a^b f(t) dt - f(x) \leq \frac{(b-x)^2}{2(b-a)} M_1$$

and if $f' \leq 0$ on $[a, b]$, then

$$(14) \quad -\frac{(b-x)^2}{2(b-a)} M_1 \leq \frac{1}{b-a} \int_a^b f(t) dt - f(x) \leq \frac{(x-a)^2}{2(b-a)} M_1.$$

Suppose now that f is a positive unimodal function on $[a, b]$, with a unique maximum $f(x_0) = M$, $x_0 \in [a, b]$. Then $f'(x) \geq 0$ for $x \in [a, x_0]$ and $f'(x) \leq 0$ for $x \in (x_0, b]$. From (2) and (3) we conclude that

$$(15) \quad 0 \geq \frac{1}{b-a} \int_a^b f(t) dt - M \geq -\frac{(x_0-a)^2 + (b-x_0)^2}{2(b-a)} M_1,$$

that is,

$$(16) \quad M(b-a) - \frac{(x_0-a)^2 + (b-x_0)^2}{2} M_1 \leq \int_a^b f(t) dt \leq M(b-a).$$

For $n = 2$, Theorem 1 yields

$$(17) \quad \left| \frac{1}{b-a} \int_a^b f(t) dt - f(x) - f'(x) \left(\frac{a+b}{2} - x \right) \right| \leq \frac{(x-a)^3 + (b-x)^3}{6(b-a)} M_2.$$

Now, for $x = (a+b)/2$ and assuming that f is a convex function, we get the following two sided HADAMARD's inequality

$$(18) \quad 0 \leq \frac{1}{b-a} \int_a^b f(t) dt - f\left(\frac{a+b}{2}\right) \leq \frac{(b-a)^2}{24} M_2$$

From (12) with $x = (a+b)/2$ and (18), we obtain the following estimate:

$$(19) \quad 0 \leq \frac{1}{b-a} \int_a^b f(t) dt - f\left(\frac{a+b}{2}\right) \leq \min\left(\frac{b-a}{4} M_1, \frac{(b-a)^2}{24} M_2\right)$$

3. AN INEQUALITY FOR THE TAIL PROBABILITY

Let X be a continuous random variable concentrated on a finite interval $[a, b]$, i.e., $P(X \notin [a, b]) = 0$, where $0 \leq a < b$. Let f be the density of X ; we assume that f is continuous on $[a, b]$ except possibly at finitely many points; let $M = \sup_{x \in [a, b]} f(x)$. Then the function $G(x) = P(X > x)$ is in $C_b^{(1)}[a, b]$, with $G'(x) = -f(x) \leq 0$, and the inequality (14) applied to the function G yields

$$(20) \quad -\frac{(b-x)^2}{2(b-a)} M \leq \frac{1}{b-a} \int_a^b P(X > t) dt - P(X > x) \leq \frac{(x-a)^2}{2(b-a)} M.$$

For a non-negative random variable X we have the following expression for the expectation:

$$E(X) = \int_0^{+\infty} P(X > t) dt,$$

hence,

$$\int_a^b P(X > t) dt = E(X) - a$$

and (20) becomes

$$(21) \quad \frac{E(X) - a - (x-a)^2 M/2}{(b-a)} \leq P(X > x) \leq \frac{E(X) - a + (b-x)^2 M/2}{(b-a)}.$$

A similar inequality has been recently obtained in [1]. We can always take $a = 0$ in (21), to obtain a simple inequality

$$(22) \quad \frac{E(X) - x^2 M/2}{b} \leq P(X > x) \leq \frac{E(X) + (b-x)^2 M/2}{b},$$

which therefore holds for any positive continuous random variable X with a piecewise continuous density, such that $P(X > b) = 0$. If X is concentrated on an interval $[-c, b]$, for some $c > 0$, then the random variable $X' = X + c$ is concentrated on $[0, b + c]$ and (22) can be applied to X' .

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